



**Monofractality of
temperatures
vs. pressure
anomalies**

A. Delière and S. Nicolay

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An inkling of the relation between the monofractality of temperatures and pressure anomalies

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Abstract

We use the discrete “wavelet transform microscope” to study the monofractal nature of surface air temperature signals of weather stations spread across Europe. This method reveals that the information obtained in this way is richer than previous works studying long range correlations in meteorological stations: the approach presented here allows to bind the Hölder exponents with the standard deviation of surface pressure anomalies, while such a link does not appear with methods previously carried out.

1 Introduction

Fractals have been extensively used in geosciences (see e.g. Arneodo et al., 2002; Blenkinsop et al., 2000; Lovejoy and Schertzer, 1995, 2013; Schertzer and Lovejoy, 1991; Schertzer et al., 2002; Tessier et al., 1993). The aim of this paper is to show that the monofractal nature of raw temperature signals is related to surface pressure anomalies. For that purpose, we first present the wavelet leaders method (WLM) as a tool for providing a multifractal formalism, which is a more recent version of the wavelet transform modulus maxima used in Arneodo et al. (2002, 1995). This method has already proven to be well-suited to study fractal objects (Abry et al., 2010; Jaffard, 2004; Jaffard and Nicolay, 2009; Lashermes et al., 2008; Wendt et al., 2009). We then use this wavelet-based approach to obtain results about the monofractality of the surface air temperature signals from a mathematical point of view. Finally, we show that the fluctuation of the monofractal exponent observed from one station to another is bonded to surface pressure anomalies. Such a relation is not observed with methods usually associated with monofractal studies previously used on these signals (Bunde and Havlin, 2002). A possible explanation could be found in the fact that the WLM can be applied to the “raw signal”.

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2 On the monofractal nature of temperature signals

Let us first recall the WLM. The discrete wavelet transform (WT) allows to decompose a signal using a single oscillating window ψ called a wavelet (Daubechies, 1992; Mallat, 1999; Meyer, 1992). The WT of a function f is defined as

$$W_\psi[f](j, k) = 2^{-j} \int f(x) \psi(2^{-j}x + k) dx,$$

where k is the space parameter and j the scale parameter (both take integer values). WT is well adapted to study the irregularities of f , even if they are masked by a smooth behavior. If f has, at a given point x_0 , a local Hölder exponent $h(x_0)$, in the sense that $|f(x) - P_{x_0}(x)| \sim |x - x_0|^{h(x_0)}$ around x_0 , where P_{x_0} is a polynomial, then with the right choice of ψ , one has $W_\psi[f](j, k) \sim 2^{-jh(x_0)}$ for the indices k such that $2^{-j}x - k$ is close to x_0 (Jaffard, 2004; Jaffard and Nicolay, 2009). The WLM is somehow a transposition of the wavelet transform modulus maxima (WTMM) to the discrete setting with a stronger theoretical background (Arneodo et al., 2002, 1995; Jaffard, 2004; Jaffard et al., 2006; Jaffard and Nicolay, 2009). Mimicking the box-counting technique, one investigates the scaling behavior of the partition function

$$S(q, j) = 2^j \sum_k (\sup_{j' \geq j} |W_\psi[f](j', k)|)^q,$$

through the function

$$\omega(q) = \lim_{j \rightarrow +\infty} \frac{\log(S(q, j))}{\log 2^{-j}},$$

where q is a real parameter. In this framework, performing a Legendre transform of ω gives a good approximation of the spectrum of singularities, defined as the Hausdorff dimension of the set of points sharing the same Hölder exponent. In other words, the

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spectrum of singularities is the function that “counts”, for a given Hölder exponent h , the number of points having h as Hölder exponent. Monofractal functions, i.e. functions with a constant Hölder exponent $h(x_0) = H$ are characterized by a linear function ω : $H = \partial\omega/\partial q$, which is equivalent to a spectrum of singularities reduced to the point $(H, 1)$. On the contrary, a nonlinear ω curve is the signature of functions displaying a multifractal behavior; in this case, h is not constant anymore and thus may fluctuate from one point to another. Let us note that the wavelet used in this work is the second order Daubechies wavelet (Daubechies, 1992), but the results remain unchanged with higher orders.

In order to confirm that the analyzed signals are monofractal, we used the “surrogate data method” (see Small et al., 2001; Theiler et al., 1992, for details). We first perform a Fourier transform of the data. Then, we randomize the Fourier phases but preserve the amplitudes and finally perform an inverse Fourier transform to create the surrogate series. Such a surrogate series has thus the same Fourier spectrum as the original data.

Since the Fourier spectrum is preserved with the surrogate procedure, the spectrum of singularities of a monofractal signal is not affected either (Daubechies, 1992; Mallat, 1999). On the other hand, if the signal is multifractal, the regularity from one point to another is modified in the surrogate data and there is thus no reason for the spectrum of singularities to be preserved. In order to illustrate this fact, we performed a test on two functions. The first is the well-known Weierstraß function defined as

$$f(x) = \sum_{j=1}^{+\infty} 2^{-j} \cos(4^j x).$$

This function is monofractal with Hölder exponent 0.5 (see Nicolay, 2006, for details). The second one is the Lebesgue–Davenport function defined as

$$f(x) = \frac{1}{2} + \sum_{n=0}^{+\infty} a_n \{2^n x\}$$

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with $a_{2n} = 2^{-n}$ and $a_{2n+1} = -2^{-n-1}$ and where $\{x\}$ is the sawtooth wave: $\{x\} = x - \lfloor x \rfloor - 0.5$ if x is not an integer and $\{x\} = 0$ else. It can be shown (see Jaffard and Nicolay, 2009, 2010) that the Lebesgue–Davenport function is multifractal with a spectrum of singularities given by $d(h) = 2h$ with $0 \leq h \leq 0.5$. The regularity of these two functions was studied with the WLM and the results are represented in Fig. 1. One can clearly see that the spectrum of singularities of a monofractal function is blind to the surrogate procedure, whereas a multifractal function and its surrogate display completely different spectra.

We applied the WLM on daily mean surface air temperature data collected from the European Climate Assessment and Dataset (<http://ecad.eu>) and computed as the mean between the daily minimum and daily maximum temperatures. In order to get homogeneous signals, we limited our study to temperature series with at least 50 years of data between 1951 and 2003 spread across Europe between 36° (Southern Spain, Italy, Greece) and 55° of latitude (Northern Ireland, Germany) and -10° (Western Ireland, Portugal) and 40° of longitude (Eastern Ukraine). By doing so, we were able to select 115 weather stations uniformly dispersed across the selected area (see Fig. 2).

For the purpose of reducing the noise, the data $f(t)$ were replaced by their temperature profiles $\sum_{u=1}^t f(u)$.

As shown in Fig. 3, one can see that the air temperature signal of Rome displays a monofractal nature: the function ω is clearly linear (coefficient of determination $R^2 > 0.995$). Also, the associated spectrum of singularities is blind to the surrogate procedure, which confirms the monofractal nature of the signal (see Fig. 4).

Such an observation holds for all the stations, which indicates that air temperature signals are monofractal. However, the value of the Hölder exponent varies from one station to another as illustrated in Fig. 3 with the stations of Rome and Armagh (Ireland). The Hölder exponents of the 115 stations range from 1.093 to 1.43, with mean 1.239 and standard deviation 0.087. These results differ from the expected values of about 0.65 (see e.g. Koscielny-Bunde et al., 1998) because we applied the WLM on the temperature profile of the raw signals for the sake of precision (as mentioned above).

These usual values (about 0.65) are recovered when the seasonal trends are removed (data not shown). Let us also remark that other methods (S^V , see Kleyntssens et al., 2015, WTMM, see Arneodo et al., 2002, 1995) based on the raw data give similar results.

5 Studies about LRC in air temperature data have been carried out using the DFA (detrended fluctuation analysis, see e.g. Bunde and Havlin, 2002; Koscielny-Bunde et al., 1998). From a methodological point of view, this method displays some similarities with the WLM. However, the DFA is concerned with LRC, not with Hölder exponents. Moreover, the DFA requires the seasonal trends to be removed whereas the
10 WLM does not; both methods are then used with the cumulative sum of the signals. Therefore, the DFA studies LRC within the summed detrended signal while the WLM allows to compute the Hölder exponent of the summed raw signal.

3 Relation with pressure anomalies: a statistical approach

15 A natural question arising is whether or not the observed Hölder exponents can be bond to some climate index. A natural choice is to try to link the surface pressure anomalies with the Hölder exponents in the following sense: can we recover the correlation structure observed in the pressure anomalies field from the spatial repartition of the Hölder exponents? Moreover, can such a structure be recovered with the DFA? To answer these questions, the map of Europe is gridded into roughly
20 200 km² pixels. We compare the map of the inverses of standard deviations of surface pressure anomalies from the NCEP-NCAR Reanalysis Project (<http://www.esrl.noaa.gov/>) with the map made of the measured Hölder exponents. From a statistical point of view, we try to show that the null hypothesis, stating that the two maps are mutually independent, can be rejected. To do so, for both maps, each pixel (corresponding to
25 an anomaly or a Hölder exponent) is normalized in order to obtain values between 0 and 1. The likeness between these maps is defined as the distance between them (considered as matrices):

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$$d = \sqrt{\sum_{i,j} (x_{i,j} - x'_{i,j})^2},$$

where $x_{i,j}$ is a pixel of the first map, $x'_{i,j}$ is the corresponding pixel of the second map and where the sum is taken over all pixels. In this case, the likeness between the maps is $d_1 = 2.68$. In order to check if these maps are akin, we use a standard Monte–Carlo method: the “Hölder map” is randomly shuffled 10 000 times. For each realization, the distance with the original anomalies map is computed in order to get a distribution of these random distances. In this way, one can look where d_1 lies in the distribution of the distances, and one can associate a p value to this particular distance d_1 . Based on the 10 000 observations, the probability $1 - p$ to have a randomly shuffled map with a distance smaller than d_1 is lower than 10^{-4} , which shows that the null hypothesis can be rejected with a high confidence level (see Fig. 5a). In other words, the higher the standard deviation of pressure anomalies, the lower the Hölder exponents.

In order to show that the DFA method carried out in Koscielny-Bunde et al. (1998) does not display analogue correlation structures, we perform the same simulation but with a map where the Hölder exponents obtained with the WLM are replaced with the values obtained with the DFA. In this case, the distance d_2 between this “DFA map” and the anomalies map is 4.68, and the probability that the distance between a randomly shuffled DFA map and the anomalies map be smaller than d_2 is $1 - p = 0.8$. This shows that the DFA map cannot be considered as correlated to pressure anomalies (see Fig. 5b). One can thus conclude that the values obtained via the DFA have no obvious relation with this climate index. Let us note that the results obtained here with the WLM do not depend on the order of the wavelet (Daubechies, 1992) and those from the DFA do not differ if the multifractal detrended fluctuation analysis (MF-DFA) is performed.

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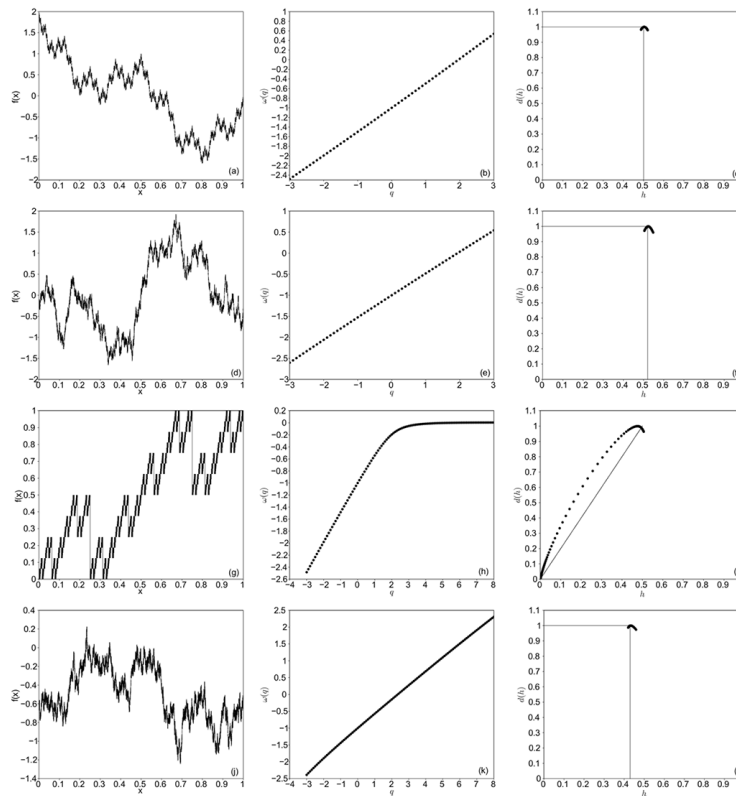


Figure 1. First column: **(a)** Weierstraß function, **(d)** a surrogate of the Weierstraß function, **(g)** Lebesgue–Davenport function, **(j)** a surrogate of the Lebesgue–Davenport function. Second and third columns represent respectively the function ω and the spectrum of singularities of the corresponding signals. One can see that the spectrum of singularities of a monofractal function is not affected by the surrogate procedure, whereas the spectrum of a multifractal function is sensitive to the procedure.

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**Figure 2.** Localization of the studied weather stations across Europe.[Title Page](#)[Abstract](#)[Introduction](#)[Conclusions](#)[References](#)[Tables](#)[Figures](#)[Back](#)[Close](#)[Full Screen / Esc](#)[Printer-friendly Version](#)[Interactive Discussion](#)

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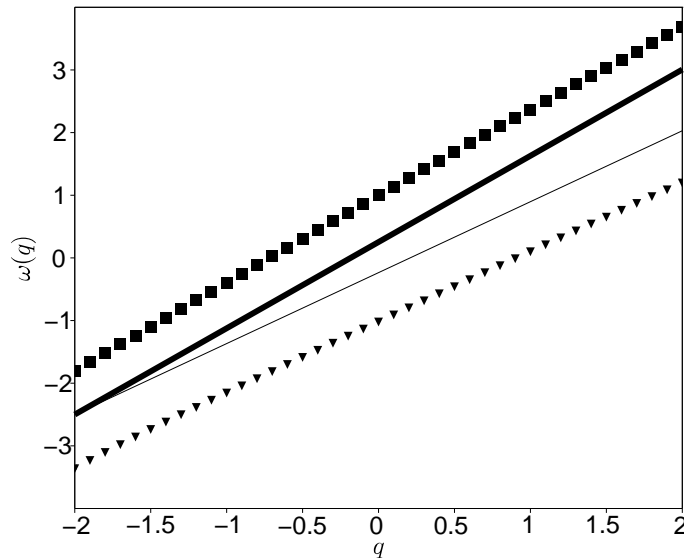


Figure 3. Comparison of the functions ω for Rome (Italy, squares) and for Armagh (Ireland, triangles). The thick straight line represents the linear regression line of ω corresponding to Rome, the other one corresponds to Armagh. Both functions ω are clearly linear, which implies that the signals are monofractal. However, since the slopes are clearly different, the associated Hölder exponents are not the same. One gets 1.4 for Rome and 1.13 for Armagh.

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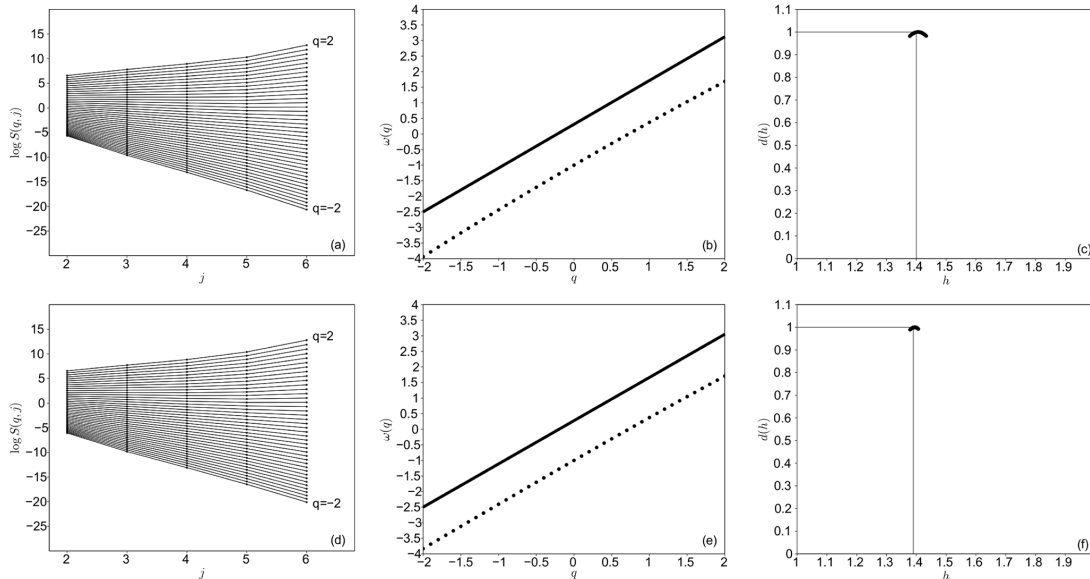


Figure 4. (a) The function $\log S(q, j)$ vs. j for q ranging from -2 to 2 (from bottom to top) by step of 0.1 for Rome. For a fixed q , the slope of the linear regression over $\log S(q, j)$ gives the value of $\omega(q)$ (see b). (b) Function ω for Rome (dots). The thick straight line represents the linear regression line of ω and shows that ω is clearly linear, which implies that the signal is monofractal with Hölder exponent given by the slope of the regression line: 1.4 . (c) Associated spectrum of singularities. Graphs (d–f) represent the corresponding figures for the surrogate signal. One can clearly see that the spectrum of singularities is not affected by the surrogate procedure, which confirms that the signal is monofractal.

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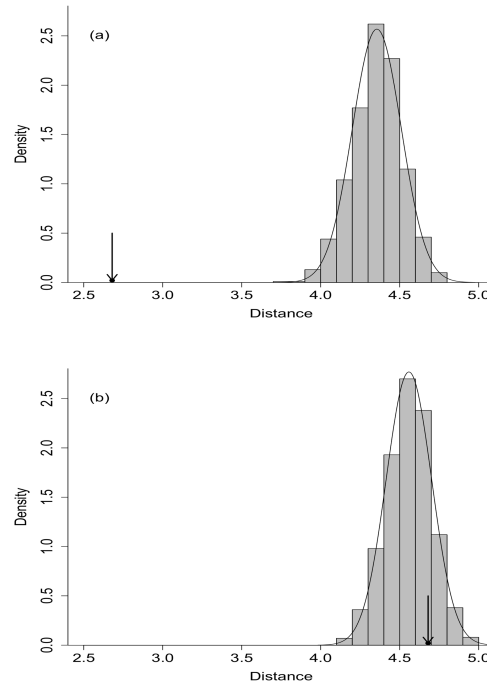


Figure 5. (a) (resp. b): histogram of the distances between the randomly shuffled Hölder maps (resp. randomly shuffled DFA maps) and the inverses of standard deviations of pressure anomalies map. The curves represent the theoretical Gaussian distributions based on the mean and the standard deviation of the measured distances. The arrow indicates the distances d_1 (resp. d_2) between the non-shuffled Hölder map (resp. DFA map) and pressure anomalies map, i.e. $d_1 = 2.68$ (resp. $d_2 = 4.68$). Obviously, original Hölder map and pressure anomalies map display similar structures, in the sense that it is extremely unlikely to measure a distance smaller than 2.68 if the Hölder map is randomly shuffled. On the opposite, exponents obtained with the DFA do not seem to be related to pressure anomalies since the distance between them is barely affected by the shuffling.

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