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On the role of storm duration in the mapping of rainfall to flood return periods

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Abstract. While the correspondence of rainfall return period T_P and flood return period T_Q is at the heart of the design storm procedure, their relationship is still poorly understood. The purpose of this paper is to shed light on the controls on this relationship. To better understand the interplay of the controlling factors we assume a simplified world with block rainfall, constant runoff coefficient and linear catchment response. We use an analytical derived flood frequency approach in which, following design practise, T_P is defined as the return period of the intensity-duration-frequency (IDF) curve given storm duration and depth. Results suggest that the main control on the mapping of rainfall to flood return periods is the ratio of storm duration and catchment response time, as would be expected. In the simple world assumed in this work, T_Q is always smaller or equal than T_P of the associated storm, i.e., $T_Q/T_P \leq 1$. This is because of the difference in the selectiveness of the rectangular filters used to construct the IDF curves and the unit hydrograph (UH) together with the fact that different rectangular filters are used when evaluating the storm return periods. The critical storm duration that maximises T_Q/T_P is, in descending importance, a function of the catchment response time and the distribution of storm duration, while the maximum value of T_Q/T_P is mainly a function of the coefficient of variation of storm duration. The study provides the basis for future analyses, where more complex cases will be examined.

1 Introduction

Flood design generally requires the estimation of flood discharges of a given return period at a site. If long stream flow records are available, the flood estimates can be derived directly from data by frequency analysis. If no or limited stream flow data are available, or floods associated with very large return periods are of interest, design floods are generally estimated based on design storms (e.g. Pilgrim and Cordery, 1993, p. 9.13). In this procedure, one or more storms of a given return period are used as an input to a rainfall-runoff model, and it is then assumed that the simulated peak discharge has the same return period (e.g. Packman and Kidd, 1980; Bradley and Potter, 1992). There have been warnings that this assumption may be grossly in error (e.g. Linsley et al., 1988, p. 365), and Pilgrim and Cordery (1975, p. 81) remarked: "The actual relationship between the frequencies of the design rainfall and the derived flood is obscure as each part of the overall design model introduces some joint probability". There are a number of "parts" that need to be considered which include storm rainfall intensity, storm duration, temporal and spatial storm patterns, and antecedent soil moisture conditions.

Despite the "obscure nature" of the mapping of return periods, there has been little work devoted to this issue. The main emphasis has been on finding the rainfall patterns or model parameters that give a close match between rainfall and flood return periods. Alfieri et al. (2008) assessed the accuracy of literature design hyetographs in producing flood peaks with the same return period as the storms. They noted that the estimation is biased in most of the cases and provided a correction factor to obtain more robust estimates of the design flood. Packman and Kidd (1980) investigated which antecedent conditions consistently gave flows that matched an observed flood frequency distribution. They found these

conditions to be related to average annual rainfall at several locations in the United Kingdom. For the simple case of the Probabilistic Rational Method, this problem has been directly addressed by estimating design runoff coefficients in a way that the design storms of a given return period produce flood peaks that match the observed flood peaks with the same return period (Pilgrim and Cordery, 1993, p. 9.20).

These studies are "black box" analyses where the emphasis is on identifying the optimum parameters in a design context. In contrast, the aim of this paper is to help understand *why* the return periods of rainfall and floods are different. Not only is this hoped to contribute to more informed flood design procedures, but this should also contribute to better assessing the magnitude of observed floods. For example, based on frequency analyses of runoff and rainfall data, Gutknecht et al. (2002) found that the 2002 Kamp flood in northern Austria had a flood return period of more than 1000 years while the return period of the associated rainfall was smaller. The difference of these magnitudes can be better understood if the controls on the relationship of the return periods are known.

As the emphasis in this paper is on *understanding* the controls, we have chosen to examine a highly simplified "world". We assume:

- block rainfall,

- constant runoff coefficient,

- linear catchment response.

This allows us to explicitly scrutinise the effect of storm duration which we consider the most basic control. The design storm procedure is mimicked in the domain of frequency distributions by an analytical derived distribution approach which facilitates the probabilistic interpretation. While the model is not complete in that it does not represent all possible processes, it is simple enough to actually understand the interplay of the controls. This paper is hence considered as a first step towards understanding more complex situations, where other processes are introduced such as the variability of the runoff coefficient, the complexity of the shapes of the hyetographs and multiple storms.

2 Design-storm procedure and definition of storm return period

The idea of the design-storm procedure is to estimate a flood of a selected return period from rainfall *intensity-duration-frequency* (IDF) *curves* for the site of interest. In many cases, the hydrological engineer has standard IDF curves available for the site but it is important to understand the procedure used to develop them. For each duration selected, the annual maximum rainfall intensity is extracted from historical rainfall records. Then frequency analysis is applied to the annual data obtaining a return period for each intensity and duration.

What is termed "duration" in the procedure is in fact not a storm duration but an aggregation time interval, or *aggregation level*. For example, if hourly rainfall data are available and one is interested in the IDF relationship for a duration of 3 h, one runs a moving averaging window over the hourly data and extracts the largest 3-h average of each year to do the frequency analysis. The moving averaging procedure is equivalent to convoluting the rainfall time series with a rectangular filter (with a base of 3 h in the example).

The way the design storm method is applied varies considerably between countries (Pilgrim and Cordery, 1993, p. 9.13) but the main components of the procedure can be summarised as following:

1. Selection of many storms of different durations reading off their mean intensities from the IDF curve corresponding to the return period T_P of interest. As noted above, rainfalls from the IDF curves do not represent complete storms but are from intense bursts within these storms. The storm duration may hence differ from the aggregation level used to read off the intensity from the IDF curve. However, in many cases storm duration is chosen equal to the aggregation level (see Chow et al., 1988, for details).

2. Application of rainfall time patterns to these storms (design hyetograph). Rigorously, the design temporal patterns need to be appropriate for the intense bursts within storms, and not for complete storms (Pilgrim and Cordery, 1993, p. 9.13) but, again, in practise these two are often set equal.

3. Application of spatial patterns to rainfall or, more simply, of an areal reduction factor for catchment area.

4. Transformation of the design storm to a flood hydrograph using a runoff model calibrated for the catchment of interest.

5. Selection of the maximum flood peak of the flood hydrographs produced by storms of different durations.

It is then assumed that this flood peak has a return period T_Q equal to T_P.

In the real world there is no rigorous solution to the problem of choosing the design parameters (i.e., the shape of the hyetograph, the rainfall-runoff model parameters, etc.) in a way that T_Q matches T_P because of the large number of controls that are difficult to understand.

In contrast, when a simplified world is assumed, the exact mapping of rainfall to flood return periods can be derived. In the case of block rainfall, as assumed here, the total rainfall event and the main burst are indeed identical, so the aggregation level used to evaluate the return period of a storm is equal to the duration of that storm.

3 The derived distribution approach in the simplified world

We use here a derived distribution approach which consists of combining a statistical rainfall model with a deterministic rainfall-runoff model. This model represents the simplified world which we are meaning to analyse. The rainfall model assumes events to be uniform and independent, and durations and intensities to be random and mutually dependent. Other factors such as multiple storms, within-storm intensity patterns, seasonality and spatial variability of the rainfall intensities are deliberately neglected.

The rainfall-runoff model is event based. It assumes a constant runoff coefficient and linear routing represented by an exponential unit hydrograph. Any non-linearities in the system as well as random controls on runoff generation are neglected. The model mimics the design storm procedure which, in this simplified world, becomes straightforward: rectangular storms, whose intensities on the IDF curve correspond to their durations, are convoluted with the linear rainfall-runoff model from which the return period of the associated flood peaks can be estimated. In this paper, we have chosen to do most of the analyses in the probability domain rather than generating time series as the former is more efficient and directly provides insight into the controls.

3.1 Derived flood return period

Given the joint probability density function of rainfall intensity i and storm duration t_r as $f_{I,T_r}(i, t_r)$, the probability that Y, i.e., the peak discharge of all independent floods, does not exceed the flood peak value q_p is

$$F_Y(q_p) = \Pr[Y \leq q_p] = \int \int_R f_{I,T_r}(i, t_r) di \, dt_r , \qquad (1)$$

where R is the region of the (i, t_r) space for which the combination of these two values is transformed into a peak smaller than or equal to q_p by the rainfall-runoff model (see Figs. 1 and 2 in Wood, 1976, for a graphical representation).

This equation can be simplified if some assumptions are formulated. We use here a simplified version of the model of Sivapalan et al. (2005). For the rainfall model we assume that the storms are independent, the number of storm events per year is Poisson distributed (see Kottegoda and Rosso, 1997, p. 455) with mean m (set to 40), and that the distribution $f_{T_r}(t_r)$ of the storm duration t_r is known.

Depending on the case analysed in this paper, we assume a discrete distribution (Sect. 4.1) or assume that t_r is continuously distributed according to a Weibull distribution (Sects. 4.2–4.4) with the probability density function (pdf):

$$f_{T_r}(t_r) = \frac{\beta_r}{\gamma_r} \left(\frac{t_r}{\gamma_r} \right)^{\beta_r - 1} \exp \left(-\frac{t_r}{\gamma_r} \right)^{\beta_r} , \qquad (2)$$

where γ_r and β_r are the scale and shape parameters, respectively. The scale parameter is related to the mean storm duration δ_r by

$$\gamma_r = \delta_r \left[\Gamma \left(1 + \frac{1}{\beta_r} \right) \right]^{-1} . \qquad (3)$$

The shape parameter is related to the coefficient of variation of the distribution by

$$CV_r = \sqrt{ \frac{\Gamma(1 + 2/\beta_r)}{[\Gamma(1 + 1/\beta_r)]^2} - 1 } . \qquad (4)$$

We examine here a basic case of $\beta_r = 0.7$ and $\delta_r = 12$ h, and vary these parameters in some of the variants, as detailed in the results section.

The storm rainfall intensity i is assumed to be distributed according to a gamma distribution (commonly used for areally averaged intensities, see e.g. Sivapalan and Blöschl, 1998):

$$f_{I|T_r}(i|t_r) = \frac{\lambda}{\Gamma(\kappa)} (\lambda i)^{\kappa - 1} \exp(-\lambda i) , \qquad (5)$$

where parameters λ and κ are functions of t_r.

Following Sivapalan et al. (2005) we assume that expectation and coefficient of variation of the intensities are related to the storm duration by power laws

$$E[i|t_r] = a_1 t_r^{b_1} \quad \text{and} \quad CV^2[i|t_r] = a_2 t_r^{b_2} , \qquad (6)$$

so that

$$\kappa = \frac{t_r^{-b_2}}{a_2} \quad \text{and} \quad \lambda = \frac{t_r^{-b_1-b_2}}{a_1 a_2} . \qquad (7)$$

In all the presented analyses, we have assumed $a_1 = 1.05$ mm h$^{-b_1-1}$, $b_1 = 0.01$, $a_2 = 1.5$ and $b_2 = -0.55$ as in Sivapalan et al. (2005) (values calibrated using data of the Frankenfels raingauge, northern Austria).

As the runoff model, we use a linear reservoir with constant response time t_c and a runoff coefficient of unity. The transformation of rainfall to runoff can then be expressed by the convolution integral of the exponential unit hydrograph (UH) with rectangular storms, from which one can evaluate the flood peak:

$$q_p = \Pi_Q(i, t_r) = i \cdot \left[1 - \exp\left(-\frac{t_r}{t_c} \right) \right] , \qquad (8)$$

where Π_Q is the runoff-model operator.

Adopting these simplifications, the integral of Equation (1) simplifies to

$$F_Y(q_p) = \int_0^\infty F_{I|T_r} \left(\Pi_Q^{-1}(q_p, t_r)|t_r \right) f_{T_r}(t_r) dt_r , \qquad (9)$$

where $\Pi_Q^{-1}(q_p, t_r) = q_p / [1 - \exp(-t_r/t_c)]$ is the storm intensity that, for given t_r and t_c, produces the flood peak q_p, and

$F_{I|T_r}(.|t_r)$ is the cumulative distribution of rainfall intensities conditioned on t_r.

The transition from the distribution of all floods $F_Y(q_p)$ to the distribution of the annual maximum floods $F_Q(q_p)$ is obtained by using the theory of order statistics (e.g. Kottegoda and Rosso, 1997, p. 455):

$$F_Q(q_p) = \exp\left\{-m\left[1 - F_Y(q_p)\right]\right\} , \tag{10}$$

which can also be expressed in terms of the return period (in years):

$$T_Q = \left\{1 - F_Q(q_p)\right\}^{-1} . \tag{11}$$

The integral in Eq. (9) is evaluated numerically in the probability space.

3.2 Derived IDF curves and storm return periods

We exploit the similarity between the runoff modelling (convolution of rainfall by a UH) and the IDF procedure (convolution of rainfall by a moving averaging window) to derive the return period of rainfall for different aggregation levels from the statistical rainfall model. The difference between the runoff model and the IDF procedure is that the former uses an exponential filter while the latter uses rectangular filters, one for each aggregation level. The exponential filter represents, to some degree, hydrological processes in that it is a summary description of the non-linearity of surface and near surface runoff (velocity depending on depth), so the rising limb is always steeper than the falling limb of the hydrograph. In contrast, the rectangular filter that is used to construct the IDF curve is commonly chosen for convenience and tradition.

If we let a random variable I denote the rainfall intensity of storms averaged on the aggregation level t_{IDF}, the probability that this intensity is lower or equal to an intensity ϕ is termed $F_I(\phi, t_{\text{IDF}})$. To derive the cumulative distribution of I (defined for a single t_{IDF}), we proceed as in Eq. (1):

$$F_I(\phi, t_{\text{IDF}}) = \Pr[I \le \phi] = \int\int_{R'} f_{I,T_r}(i, t_r)\,di\,dt_r , \tag{12}$$

where R' is the region of the (i, t_r) space such that the combination of these two values is transformed to a value smaller or equal to ϕ by the IDF filter with aggregation level t_{IDF}. The result of the rectangular filtering can be written as:

$$\phi = \Pi_P(i, t_r) = \begin{cases} i & \text{if } t_{\text{IDF}} \le t_r \\ i \cdot t_r / t_{\text{IDF}} & \text{if } t_{\text{IDF}} > t_r \end{cases} . \tag{13}$$

With analogous simplifications as in Eq. (9), the cumulative distribution of I is

$$F_I(\phi, t_{\text{IDF}}) = \int_0^\infty F_{I|T_r}\left(\Pi_P^{-1}(\phi, t_r)|t_r\right) f_{T_r}(t_r)\,dt_r , \tag{14}$$

where $\Pi_P^{-1}(\phi, t_r)$ is the inverse function of Eq. (13) and expresses the intensity of a storm of duration t_r that has average
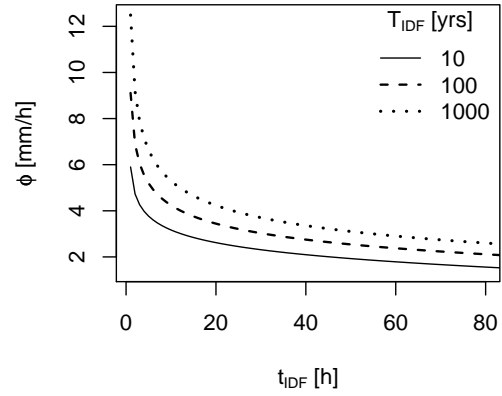
Fig. 1. Example IDF curves with parameters a_1=1.05, b_1=0.01, a_2=1.5, b_2=−0.55, β_r=0.7 and δ_r=12 h of the rainfall model. ϕ is the rainfall intensity over the aggregation level t_{IDF} and T_{IDF} is the IDF return period.

intensity ϕ over the aggregation level t_{IDF}. The transition from the distribution $F_I(\phi, t_{\text{IDF}})$ of all intensities over one aggregation level t_{IDF} to the distribution $F_P(\phi, t_{\text{IDF}})$ of the annual maximum rainfall intensities P yields

$$F_P(\phi, t_{\text{IDF}}) = \exp\left\{-m\left[1 - F_I(\phi, t_{\text{IDF}})\right]\right\} , \tag{15}$$

which can also be expressed in terms of the return period (in years) as

$$T_{\text{IDF}}(\phi, t_{\text{IDF}}) = \left\{1 - F_P(\phi, t_{\text{IDF}})\right\}^{-1} . \tag{16}$$

This equation represents the IDF curves, an example of which is provided in Fig. 1.

In our simplified world, the return period of individual storms can now be read off the IDF curve as $T_{\text{IDF}}(\phi=i,\ t_{\text{IDF}}=t_r)$. The return period T_P of the storms that produce the maximum annual peaks q_p (here called *flood-producing storms*) is then

$$T_P = T_{\text{IDF}}(\phi = \Pi_Q^{-1}(q_p, t_r = t_{\text{IDF}}),\ t_{\text{IDF}} = t_r) \tag{17}$$

where $\Pi_Q^{-1}(.)$ is the storm intensity that, for given t_r and t_c, produces the flood peak q_p.

4 Results

The mapping of rainfall to flood return periods is now shown by graphs that relate the return periods T_Q and T_P of the same event. Different storm durations are drawn as different lines in the (T_P, T_Q) space. The envelope of these lines, which maximises the ratio T_Q/T_P, is the result of the design storm procedure described in Sect. 2. We term critical storm duration(s), t_r^*, the duration(s) for which T_Q/T_P is maximised.
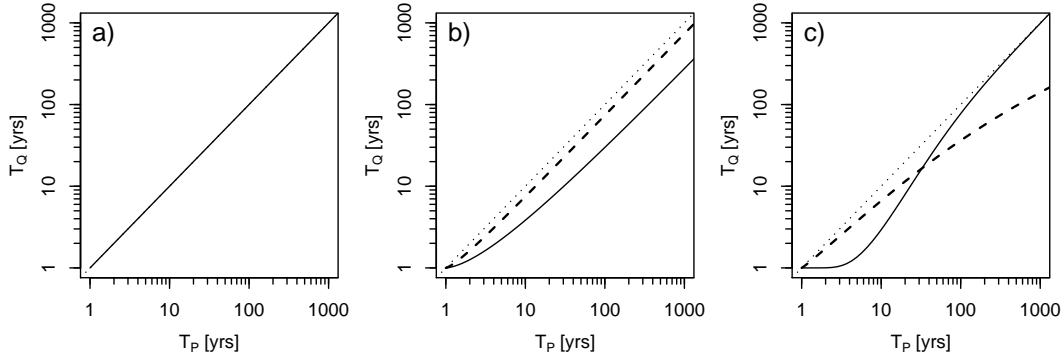
Fig. 2. Relation between rainfall return periods T_P and flood return periods T_Q for discrete distributions of storm durations: **(a)** only one possible storm duration t_1; **(b)** only two possible storm durations $t_1=6$ h (continuous line) and $t_2=12$ h (dashed line); **(c)** only two possible storm durations $t_1=3$ h (continuous line) and $t_2=144$ h (dashed line). The response time of the catchment is always $t_c=6$ h.

4.1 Simple case: discrete number of storm durations

First, we consider some extremely simplified cases. Suppose to have only one possible storm duration t_1, so that its probability is $p_{T_r}(t_1)=1$. In this limiting case, as storms differ by their intensity only, q_p is a strictly monotonically increasing function of i and a perfect correspondence $T_P=T_Q$ follows (Fig. 2a).

If, instead, two storm durations t_1 and t_2 are possible with equal probabilities $p_{T_r}(t_1)=p_{T_r}(t_2)=1/2$, this is no longer true. In Fig. 2b we have imposed $t_1=t_c$ and $t_2=2t_c$. The graph suggests that T_Q is always lower than T_P. This result can be explained by the differences between the distribution of the flood peak q_p and the distributions of the intensities that can cause it, i.e., $\Pi_Q^{-1}(q_p,t_1)$ and $\Pi_Q^{-1}(q_p,t_2)$. From Equation (9) it follows that

$$F_Y(q_p) = \frac{1}{2}\left[F_{I|T_r}\left(\Pi_Q^{-1}(q_p,t_1)|t_1\right) + \right.$$
$$\left. + F_{I|T_r}\left(\Pi_Q^{-1}(q_p,t_2)|t_2\right)\right] . \quad (18)$$

On the other hand, since two durations are possible, two distributions F_I have to be considered. For the storms of duration t_1 (continuous line in Fig. 2b), given that

$$\Pi_P^{-1}\left(\Pi_Q^{-1}(q_p,t_1),t_1\right)=\Pi_P^{-1}\left(\Pi_Q^{-1}(q_p,t_1),t_2\right)=\Pi_Q^{-1}(q_p,t_1), \quad (19)$$

Equation (14) becomes

$$F_I\left(\Pi_Q^{-1}(q_p,t_1),t_1\right) = \frac{1}{2}\left[F_{I|T_r}\left(\Pi_Q^{-1}(q_p,t_1)|t_1\right) + \right.$$
$$\left. + F_{I|T_r}\left(\Pi_Q^{-1}(q_p,t_1)|t_2\right)\right] . \quad (20)$$

The only difference between F_Y and F_I is in the second part of the second term of Eqs. (18) and (20). As $t_1<t_2$, it follows that $F_{I|T_r}(\Pi_Q^{-1}(q_p,t_1)|t_2)$ is greater than $F_{I|T_r}(\Pi_Q^{-1}(q_p,t_2)|t_2)$ and, as a consequence, that $F_I(\Pi_Q^{-1}(q_p,t_1),t_1)>F_Y(q_p)$. Since they transform to the extremes in the same way, $T_P>T_Q$.

Considering the storms of duration t_2 (dashed line in Fig. 2b) Eq. (14) becomes

$$F_I\left(\Pi_Q^{-1}(q_p,t_2),t_2\right) = \frac{1}{2}\left[F_{I|T_r}\left(\Pi_Q^{-1}(q_p,t_2)\cdot\frac{t_2}{t_1}|t_1\right) + \right.$$
$$\left. + F_{I|T_r}\left(\Pi_Q^{-1}(q_p,t_2)|t_2\right)\right] . \quad (21)$$

Here the difference between F_Y and F_I is in the first part of the second term of Eqs. (18) and (21). As $t_1<t_2$, it follows that $F_I(\Pi_Q^{-1}(q_p,t_2),t_2)>F_Y(q_p)$, so $T_P>T_Q$.

Heuristically, the finding of T_Q always being lower than T_P has to do with (a) the shape of the UH being different from the shape of the IDF filter and (b) the fact that two different filters are used for assigning T_P while a single filter is used for assigning T_Q. The rectangular IDF filter of aggregation level t_1 strongly highlights the events of duration t_1, more than the exponential UH does. In the same way, the IDF filter of aggregation level t_2 strongly highlights the events of duration t_2. The UH is less selective and easily allows events of both durations to produce maximum peaks.

Another system is shown in Fig. 2c that differs from 2b in that $t_1\ll t_c$ and $t_2\gg t_c$. In this case, for high return periods T_Q is similar to T_P of short storms (duration t_1), for low return periods T_Q is similar to T_P of long storms (duration t_2), while in the middle part T_Q is lower than T_P for both types of storms. This behaviour is a consequence of the shape of $f_{I|T_r}(i|t_r)$, i.e., the distribution of intensities given the storm durations. Consider once more the case $t_{IDF}=t_1$ (continuous line in Fig. 2c) for which Eqs. (18) and (20) hold. For high values of q_p, the terms $F_{I|T_r}(\Pi_Q^{-1}(q_p,t_1)|t_2)$ and $F_{I|T_r}(\Pi_Q^{-1}(q_p,t_2)|t_2)$, that are the ones that cause the differences between T_P and T_Q, tend to be equal to 1. Therefore $T_Q\approx T_P$. Analogous considerations can be used to explain the behaviour of the dashed line corresponding to $t_{IDF}=t_2$. For low values of q_p, the terms $F_{I|T_r}(\Pi_Q^{-1}(q_p,t_1)|t_1)$ and $F_{I|T_r}(\Pi_Q^{-1}(q_p,t_2)\cdot\frac{t_2}{t_1}|t_1)$ tend both to 0.

Hydrologically, the most extreme short duration storms result in relatively large floods (larger than those from the most
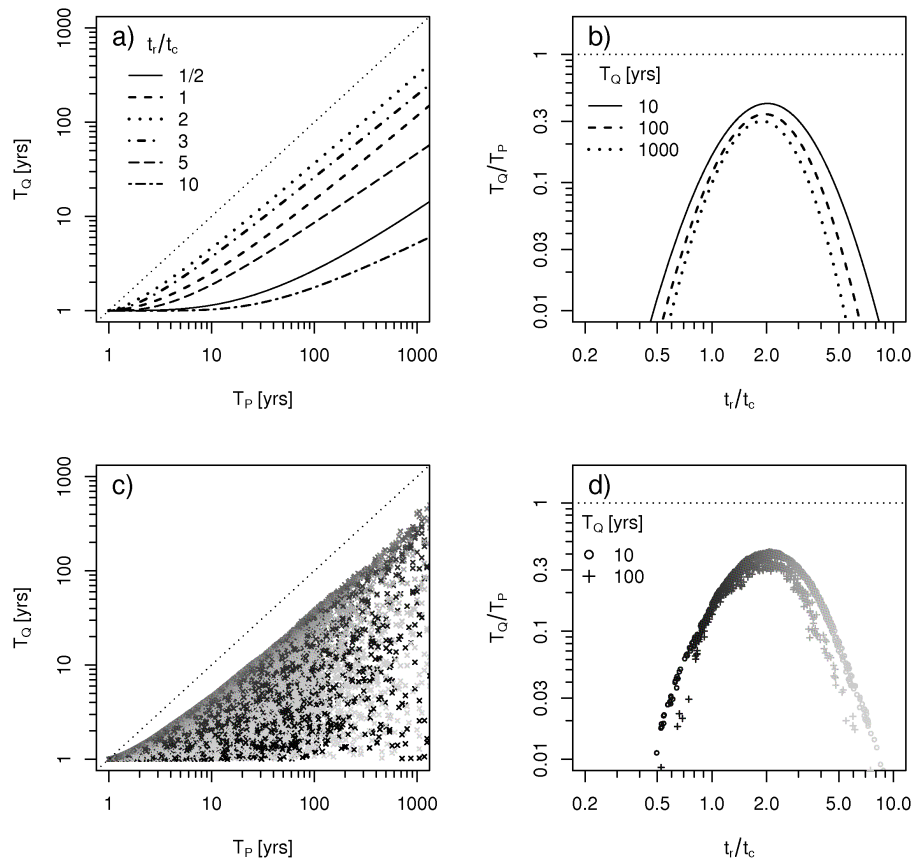
Fig. 3. Relation between rainfall return periods T_P and flood return periods T_Q for different storm durations t_r. Case of response time t_c equal to the mean storm duration δ_r ($t_c = \delta_r = 12$ h). The two upper graphs **(a, b)** are obtained with the analytical derivation of T_P and T_Q; the two lower graphs **(c, d)** are the result of a Monte-Carlo simulation (20 000 simulated years). The grey-scale of points represents storm durations (dark points are short storms). For the slices in graph (d), flood return periods between 8 and 13 years have been labelled as 10 yrs and those between 50 and 200 years have been labelled as 100 yrs.

extreme long duration storms) because of the shape of the IDF curve. The smallest long duration storms result in relatively large floods because saturation of the catchment is reached.

In the case of Fig. 2a the design storm procedure will give a 1:1 mapping of the return periods. In the case of Fig. 2b, t_2 will be the critical storm duration ($t_r^* = t_2$). It is interesting to see that in Fig. 2c the highest maximum annual floods are given by rainfall events with different durations, depending on the storm return period considered, so t_r^* is not unique for the catchment.

Analogous results can be obtained when 3 or more storm durations are possible. The generalised case is given by the continuous distribution of storm durations shown in the next section.

4.2 Continuous distribution of storm durations

Storm duration t_r is now considered to be distributed according to the Weibull distribution of Eq. (2). Figure 3 shows the relationship between T_P and T_Q for different storm durations

t_r when the response time of the linear reservoir t_c is equal to the average annual storm duration δ_r. The mapping of the return periods is plotted in the upper-left graph. The six curves represent 6 storm durations as multiples of t_c.

The envelope curve in Fig. 3a is always below the 1 to 1 line indicating that T_Q obtained by the design storm procedure is smaller than T_P. This can be explained analytically by generalizing the comparison between Eqs. (18), (20) and (21) to the continuous case, and analogous hydrological arguments apply as in the discrete case. T_P is greater than T_Q because the IDF filters are more selective than the UH filter. This is because of their shapes (rectangular as opposed to exponential) and the fact that, for each storm duration, a different filter is used when constructing the IDF curve while the same UH is used for all storms.

This behaviour is further illustrated by slicing Fig. 3a horizontally, and expressing the slices in terms of the ratio of return periods T_Q/T_P. This ratio is shown in Fig. 3b as a function of storm duration t_r for three flood return periods T_Q (10, 100 and 1000 years). These curves are obtained by
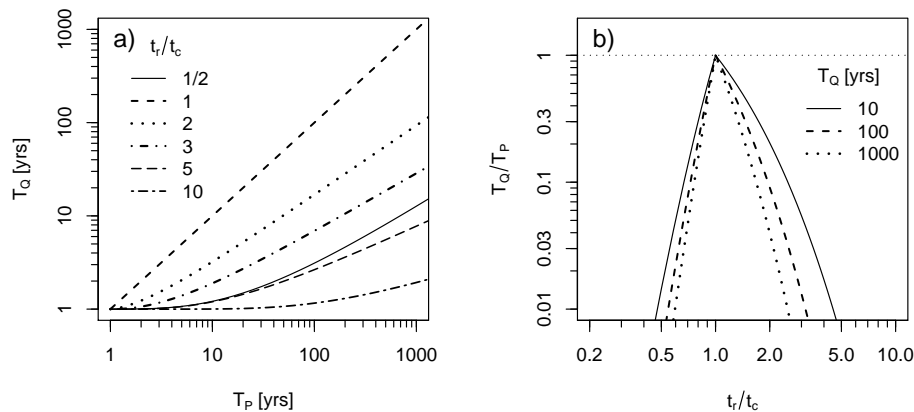
Fig. 4. Relation between rainfall return periods T_P and flood return periods T_Q for different storm durations t_r. Case of rectangular UH with width equal to t_c ($t_c = \delta_r = 12$ h).

interpolations in the (T_P, T_Q, t_r) space. For short storm durations, the ratio of the return periods increases with storm duration, reaches a maximum around $t_r^* \approx 2t_c$ (the critical storm duration), and decreases for longer durations. At the maximum, the difference between T_P and T_Q slightly depends on the considered value of T_Q, and T_Q/T_P is always below 0.5.

It appears that, for short durations, peak flows are small and hence the ratio T_Q/T_P is small. During short duration storms, the catchment does not contribute entirely to runoff generation, so the storms must be extremely severe to produce big floods. The ratio at short durations is therefore controlled by runoff response. For long durations, peak flows are also small and hence the ratio T_Q/T_P is small. This is because, in this case, runoff has already reached equilibrium (the catchment totally contributes to runoff) but, according to the IDF curves, rainfall intensities are very low. The ratio at long durations is therefore controlled by rainfall. The maximum of the curves in Fig. 3b is due to the interplay between catchment processes and rainfall processes and occurs at the critical storm duration t_r^*.

To be able to plot individual events, Monte-Carlo simulations have been performed based on the same assumptions for the rainfall and rainfall-runoff models. The results are hence fully consistent with the analytical approach. A scatter plot of T_P versus T_Q based on the Monte-Carlo simulations is shown in Fig. 3c. Each point represents one maximum annual flood and its corresponding storm. The scatter comes from random storm durations. The grey scale used in the graph is related to the duration of the storms: dark points are short storms, light-grey points are long storms. Two slices through Fig. 3c represented in terms of the ratio T_Q/T_P are shown in Fig. 3d. The points, each one representing one event, clearly trace the analytical curves of Fig. 3b. As can be seen in the graphs, the simulated events are more dense around the critical storm duration (corresponding to the envelope curve in Fig. 3c), meaning that the catchment acts as a

filter on storm durations. This is an important effect that cannot be obtained by the analytical approach, and is discussed in more detail in Sect. 4.4.

In Sect. 4.1 we have stressed that the shape of the UH is an important control on the relation $\max(T_Q/T_P) < 1$. We hence expect the behaviour of the mapping to change when changing the shapes of the UH. In the limiting case, where the UH is a rectangular function of width t_c, the resulting mapping is shown in Fig. 4. In this case, the critical storm duration is $t_r^* = t_c$ and gives $T_P = T_Q$. This is not surprising, because the convolution of the rainfall series with the rectangular UH is identical to the IDF filtering. The series of maximum peaks q_p and the series of the peak intensities ϕ (with $t_{IDF} = t_c$) coincide. For shorter and longer storms than t_c, the flood return periods are of course smaller than the rainfall return period.

For the case of a rectangular UH, the design storm method would hence be exact yielding the flood return period equal to the rainfall return period. UH shapes that lie between exponential and rectangular have been represented by a beta distribution, as in Alfieri et al. (2008). Results (not shown here) indicate that the envelope curves lie between those represented in Figs. 3a and 4a.

4.3 Effect of average storm duration, catchment response time and return period

After these first examples, we investigate the combined effect on the (T_P, T_Q) mapping of the average storm duration δ_r, the catchment response time t_c and the range of return periods T considered. Figure 5 is analogous to Fig. 3ab except for $t_c > \delta_r$ ($t_c = 60$ h, $\delta_r = 12$ h). In this case, the critical storm duration visibly depends on the return period which was not the case for $t_c = \delta_r = 12$ h (Fig. 3). For low return periods, $t_r^* \approx t_c$, while for high return periods $t_r^* \approx 2t_c$.

This is the effect of the interplay of the catchment response time t_c and the distribution of storm durations. Almost all storms are short (relative to t_c), so regular floods are produced by short storms. Only very extreme floods are
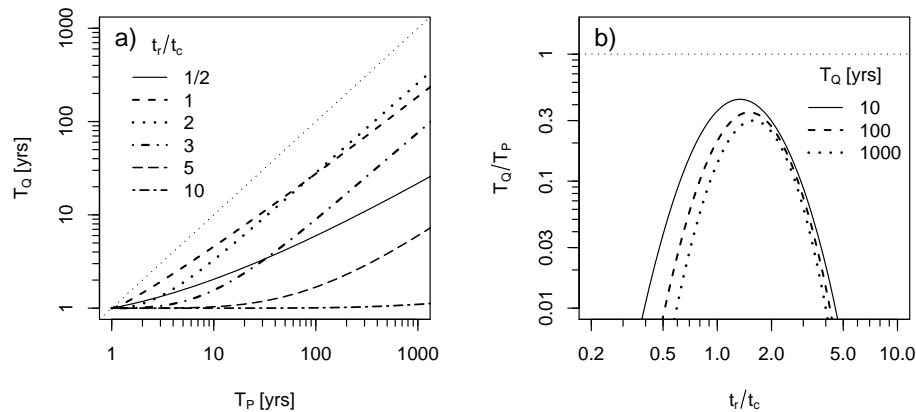
Fig. 5. Relation between rainfall return periods T_P and flood return periods T_Q for different storm durations t_r. Case of exponential UH with response time t_c greater than the mean storm duration δ_r (t_c=60 h, δ_r=12 h).

produced by (rare) longer duration storms resulting in larger t_r^* as the return period increases. Interestingly, for the converse case (t_c=60, δ_r=12 h) t_r^* decreases from about $3t_c$ to about $2t_c$ as the return period increases (not shown here). Apparently, for the limit of very large return periods, the limiting critical storm duration t_r^* is about $2t_c$. This can be interpreted as follows. For large return periods, the maximum tends to approach the time of concentration (for reasons consistent with the rational method) as there is a large number of storms that may produce that flood, so the distribution of storm durations becomes relatively less important. The time of concentration (when the water from the furthest point of the catchment reaches the outlet) is, typically, larger than the response time (e.g. defined as the recession parameter), which may explain the maximum at twice the response time ($2t_c$).

It is clear that the distribution of possible storm durations is an important control on the mapping. Figure 6a and b are generalisations of Figs. 3 and 5 and show the effects of changing the mean storm duration. In Fig. 6a only the envelope curves are shown for t_c=12 h while δ_r varies. The figure indicates that the average duration does not affect the position of this curve.

In Fig. 6b the lines correspond to a return period of T_Q=100 years. The figure indicates that as δ_r increases, the maximum of the curves moves to the right, i.e., the ratio t_r^*/t_c increases as well. This means that the critical storm duration depends both on catchment processes (here parameterised by t_c) and on rainfall processes (δ_r). It is interesting that the critical storm duration changes with the distribution of durations but the envelope curve max(T_Q/T_P) does not.

In Fig. 6c and d, the coefficient of variation CV_r of the distribution of t_r has been changed by varying the shape parameter β_r of the Weibull distribution (see Eq. 4). As CV_r tends to 0, the flood return period approaches the rainfall return period (i.e., a 1 to 1 mapping) as only a single storm duration is possible. This is the case of Fig. 2a. On the other hand, if CV_r increases, max(T_Q/T_P) decreases and reaches

a stable value of about 0.3 for very dispersed distributions. This is because, for high CV_r, it is mainly the tail of the distribution that changes. For the durations commensurate with the catchment response time t_c, the changes are small.

4.4 The catchment as a filter on storm durations

In the previous sections we have derived flood and storm return periods for different storm durations. For example, considering Fig. 3b, a storm of duration $t_r \approx t_c$ would produce a flood of return period $T_Q \approx 10$ years if $T_P \approx 50$ years, provided such as storm occurs. Figure 3a and b gives no information about the probability of storm occurrence. It is now of interest to understand what is the probability that a maximum annual flood was produced by an event of duration t_r. The distribution of this probability is called here *distribution of flood-producing storm durations*, $f_{T_r}^*$. Some cases obtained by Monte-Carlo simulations are shown in Fig. 7. In both graphs, the Weibull parent distribution of t_r of Eq. (2) is shown for comparison (solid line). The transition to the extremes transforms it into the dashed and dashed-dotted lines. As shown in Fig. 7a, the distribution of flood-producing storm durations depends on the catchment response time t_c (here the exponential UH is used). Given a "parent" distribution of storm durations (e.g. Eq. 2), the catchment acts as a filter (see also Skøien and Blöschl, 2006). Storms with durations on the order of magnitude of the catchment response time are transformed into annual floods with higher probability. If t_c increases, also the average and the spread of the distribution of flood-producing t_r increase. In small catchments with short response times, it is usually the short duration storms that produce large floods (e.g. flash-floods), and the duration of the storms does not vary much. On the other hand, in large catchments with long response times, long duration storms tend to produce large floods and there is a larger variability of flood-producing storm durations. Short flash floods, however, rarely produce large floods in large catchments. Clearly, this is the conceptual basis of the rational method.
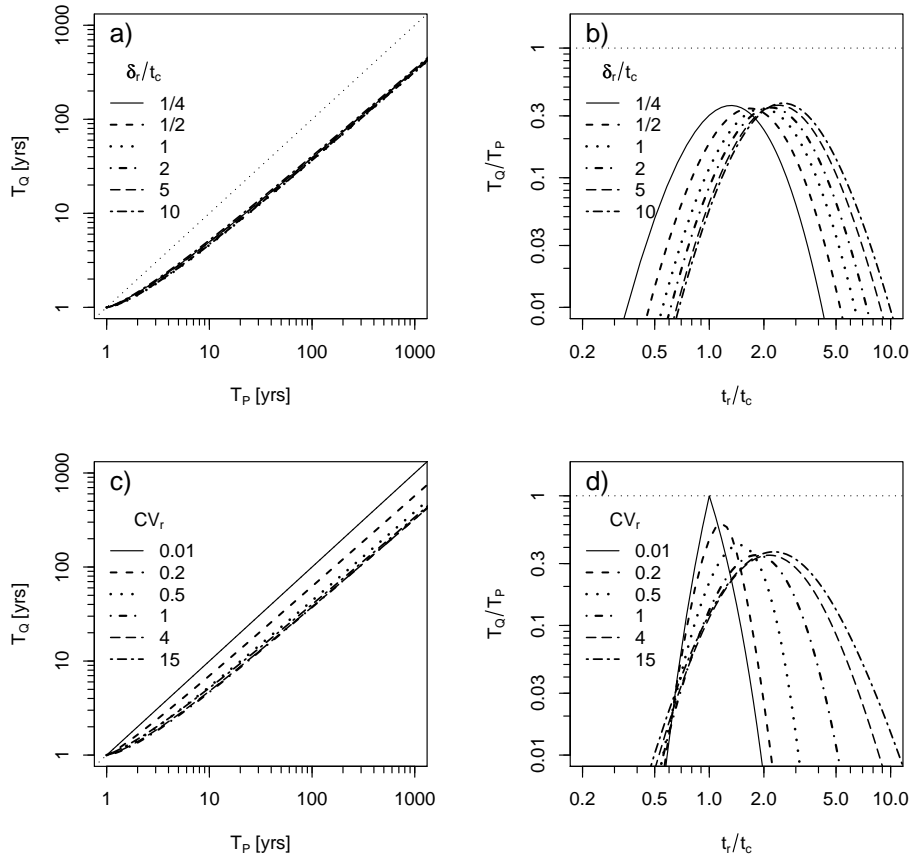
Fig. 6. Relation between rainfall return periods T_P and flood return periods T_Q for different distributions of the storm duration t_r: **(a, b)** different relations between catchment response time t_c and mean storm duration δ_r when the coefficient of variation $CV_r \approx 1.46$ ($\beta_r = 0.7$) and $t_c = 12$ h; **(c, d)** different CV_r when $\delta_r = t_c = 12$ h; (a, c) envelope curves, result of the design storm procedure; (b, d) horizontal slices for $T_Q = 100$ years.
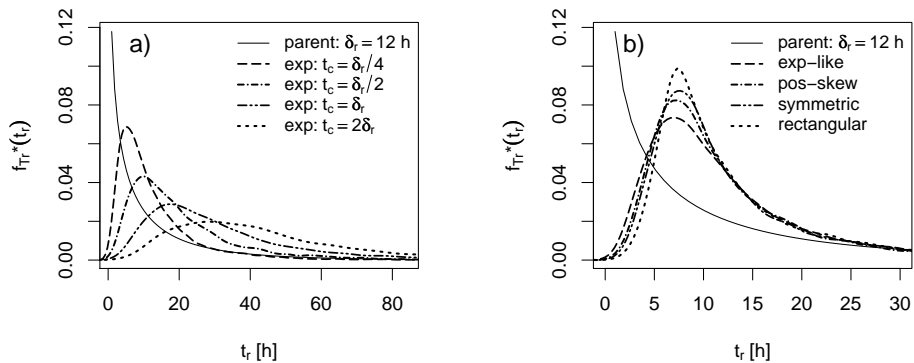
Fig. 7. Effect of the response time t_c **(a)** and the shape of the UH **(b)** on the flood-producing storm durations. The black continuous line is the parent distribution, given in Eq. (2) with mean storm duration $\delta_r = 12$ h and coefficient of variation $CV_r \approx 1.46$ ($\beta_r = 0.7$), while the other lines are the distributions of the durations of flood-producing storms. In panel (a) the UH is a linear reservoir (see Eq. 8) with response time $t_c = 3, 6, 12, 24$ h; in panel (b) the UHs have different shapes given by beta distributions with shape parameters p_β and q_β: exponential like ($p_\beta = 1$, $q_\beta = 4$, width of 17 h); positively skewed distribution ($p_\beta = 2$, $q_\beta = 4$, width of 12 h); symmetric distribution ($p_\beta = 2$, $q_\beta = 2$, width of 9 h); rectangular ($p_\beta = 1$, $q_\beta = 1$, width of 7 h).
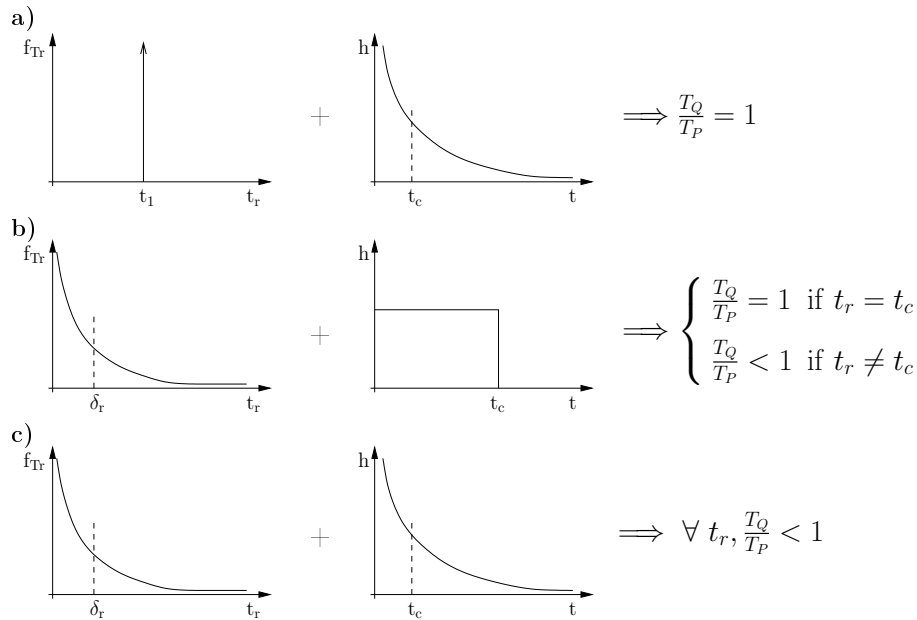
Fig. 8. Schematic summary of the results: **(a)** one possible storm duration t_1 and exponential UH; **(b)** Weibull distribution of the storm durations t_r and rectangular UH; **(c)** Weibull distribution of the storm durations t_r and exponential UH.

Figure 7b shows instead the effects of the UH shape. We have considered five cases using a beta distribution for the UH with shape parameters p_β and q_β as in Alfieri et al. (2008). When $p_\beta=1$ and $q_\beta=4$, the UH is similar to the exponential one; when $p_\beta=2$ and $q_\beta=4$, it peaks at about 1/4 of the UH width; when $p_\beta=2$ and $q_\beta=2$, it is symmetric with the maximum in the center; and when $p_\beta=1$ and $q_\beta=1$, it is a rectangular UH. The widths of the beta UHs have been chosen in order to have a similar mode of the distribution of flood-producing storm durations (about 7 h), which implies similar response times.

The results show that the distribution of the durations of the flood-producing storms associated with the rectangular UH has the largest peak of all the distributions. This means that the rectangular filter is more selective on storm durations than the others. The exponential-like UH, in contrast, is the least selective. The difference in selectiveness of the exponential and rectangular filters is at the heart of the differences between the rainfall and flood return periods together with the fact that different rectangular filters are used for each t_{IDF} when evaluating the storm return periods.

5   Conclusions

The purpose of this research is to contribute to a better understanding of the design storm method used in engineering hydrology and a better understanding of the relationship between storm and flood return periods of observed floods. This relationship is controlled by many factors of which we examine storm duration as a starting point. Based on an an-

alytical approach we derive the extreme storm and flood frequency distributions from a rainfall model assuming block rainfall, constant runoff coefficient and linear catchment response.

Even in the simple world modelled here, the relationship between storm and flood return periods is not straightforward. The main findings are summarised in Fig. 8:

a) If there is only a single storm duration, the flood return period T_Q is always equal to the rainfall return period T_P, irrespective of the shape of the unit hydrograph.

b) In a more realistic case where storm durations vary, the flood and rainfall return periods are still identical, provided runoff response can be represented by a rectangular unit hydrograph and the storm duration is equal to the duration of the unit hydrograph. For longer and shorter storms $T_Q < T_P$.

c) In the more general case of a non-rectangular right skewed UH and varying storm durations, $T_Q < T_P$ always. This is because of the difference in the selectiveness of the rectangular filter used to construct the IDF curve and the UH, together with the fact that different rectangular filters are used for each t_{IDF} when evaluating the storm return periods.

A critical storm duration t_r^* exists where T_Q/T_P is at a maximum. For case (c), t_r^* is (in decreasing importance) a function of the catchment response time t_c, average storm duration δ_r, the coefficient of variation of the storm duration CV_r and the rainfall return period T_P. t_r^* ranges between t_c

and $3t_c$, mainly depending on δ_r. The maximum of T_Q/T_P (when varying storm duration) is about 0.4 implying that, for the simple world examined here, the design storm method would give a 40 year flood when using a 100 year storm as an input to the runoff model.

We also examined the duration of the storms that produce the maximum annual floods. As the catchment response time t_c increases, so does the average and the spread of the distribution of flood-producing storm durations. This analysis also confirmed the higher selectiveness of the rectangular filters used to calculate the IDF curves as compared to the exponential unit hydrograph.

Comprehensive sensitivity analyses (not shown in this paper) indicate that the above results are generic and do not depend much on the particular rainfall model used. For a world where

- storm duration varies,

- rainfall intensities are distributed according to a positively skewed distribution,

- extreme rainfall intensity decreases with storm duration

and considering the simplifying assumptions made in this paper

- block rainfall

- constant runoff coefficient

- linear catchment response

the mapping of rainfall to flood return periods will always look very similar to the results shown here.

While these results are only applicable to this simple world, they are an important stepping stone towards the more complex cases involving random runoff coefficients and complex hyetographs, for example. For these, one would expect that the envelope curve in the (T_P, T_Q) space goes above the 1 to 1 line as is observed for floods that have occurred (e.g. Gutknecht et al., 2002). These additional factors will be examined in a follow up paper.

Notation

Variables:

T_Q flood return period [yrs]

T_P storm return period [yrs]

T_{IDF} IDF return period [yrs]

t_r, t_1, t_2 storm duration [h]

i average storm intensity [mm/h] over its duration t_r

t_{IDF} aggregation level [h]

ϕ average storm intensity [mm/h] over the aggregation level t_{IDF}

q_p flood peak [mm/h]

t_r^* critical storm duration [h]

R, R' regions of the (i, t_r) space

Y, Q, I, P random variables

Parameters (values for the basic case are indicated):

δ_r mean storm duration (12 h)

CV_r coefficient of variation of the storm duration (1.46)

β_r Weibull shape parameter of storm duration (0.7)

γ_r Weibull scale parameter of storm duration (9.48 h)

λ Gamma scale parameter of storm intensity

κ Gamma shape parameter of storm intensity

a_1, b_1, a_2, b_2 rainfall model parameters ($a_1 = 1.05 \, \text{mm h}^{-b_1-1}$, $b_1 = 0.01$, $a_2 = 1.5$, $b_2 = -0.55$)

m average number of storms per year (40)

p_β, q_β shape parameters of the Beta distribution

Operators:

$E(.)$ expected value

$CV(.)$ coefficient of variation

$F_X(.)$ cumulative distribution function of X

$f_X(.)$ probability density function of X

$p_X(.)$ probability mass function of X

$f_{X,Y}(.)$ joint probability density function of X and Y

$f_{X|Y}(.)$ conditional probability density function of X given Y

$\text{Pr}(.)$ probability

$\Gamma(.)$ Gamma function

$\Pi_Q(.)$ runoff model operator

$\Pi_P(.)$ IDF filter operator

References

Alfieri, L., Laio, F., and Claps, P.: A simulation experiment for optimal design hyetograph selection, Hydrol. Process., 22(6), 813–820, doi:10.1002/hyp.6646, 2008.

Bradley, A. A. and Potter, K. W.: Flood frequency analysis of simulated flows, Water Resour. Res., 28, 2375–2385, 1992.

Chow, V. T., Maidment, D. R., and Mays, L. W.: Applied Hydrology, Civil Engineering Series, McGraw-Hill Book Company, International Edn., 572 pp., 1988.

Gutknecht, D., Reszler, C., and Blöschl, G.: Das Katastrophenhochwasser vom 7. August 2002 am Kamp – eine erste Einschätzung (The August 7, 2002 – flood of the Kamp – a first assessment), Elektrotechnik und Informationstechnik, 119, 411–413, 2002.

Kottegoda, N. T. and Rosso, R.: Statistics, Probability, and Reliability for Civil and Environmental Engineers, McGraw-Hill Companies, International Edn., 735 pp., 1997.

Linsley, R. K., Kohler, M. A., and Paulhus, J. L. H.: Hydrology for Engineers, McGraw-Hill Book Company, London, Si Metric Edn., 492 pp., 1988.

Packman, J. C. and Kidd, C. H. R.: A logical approach to the design storm concept, Water Resour. Res., 16, 994–1000, 1980.

Pilgrim, D. H. and Cordery, I.: Rainfall temporal patterns for design floods, J. Hydr. Eng. Div.-ASCE, 101, 81–95, 1975.

Pilgrim, D. H. and Cordery, I.: Flood Runoff, in: HandBook of Hydrology, edited by Maidment, D. R., McGraw-Hill Companies, International Edn., Chap. 9, 42 pp., 1993.

Sivapalan, M. and Blöschl, G.: Transformation of point rainfall to areal rainfall: Intensity-duration frequency curves, J. Hydrol., 204, 150–167, 1998.

Sivapalan, M., Blöschl, G., Merz, R., and Gutknecht, D.: Linking flood frequency to long-term water balance: Incorporating effects of seasonality, Water Resour. Res., 41, W06012, doi: 10.1029/2004WR003439, 2005.

Skøien, J. O. and Blöschl, G.: Catchments as space-time filters – a joint spatio-temporal geostatistical analysis of runoff and precipitation, Hydrol. Earth Syst. Sci., 10, 645–662, 2006, http://www.hydrol-earth-syst-sci.net/10/645/2006/.

Wood, E. F.: An analysis of the effects of parameter uncertainty in deterministic hydrologic models, Water Resour. Res., 12, 925–932, 1976.